

today: finish prob. review
parametric models

independence:

X is independent of $Y \iff p(x,y) = p(x)p(y) \quad \forall (x,y) \in \Omega_x \times \Omega_y$

notation: $X \perp\!\!\!\perp Y$

• R.V. X_1, \dots, X_n are "mutually independent" $\iff p(x_{1:n}) = \prod_{i=1}^n p(x_i)$
 $\forall x_{1:n} \in \prod_{i=1}^n \Omega_{x_i}$

side note: $Z = (X, Y)$
 cts. discrete

$$P\{X \in \text{box}, Y=y\} = \int_{x \in \text{box}} \overset{\text{pdf part}}{p(x,y)} dx$$

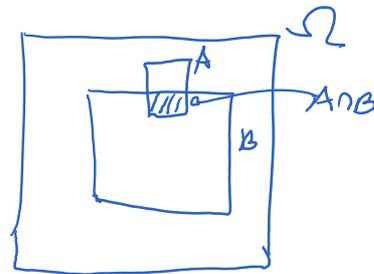
$$P\{X \in \text{box}, Y \in \text{set}\} = \int_{x \in \text{box}} \left(\sum_{y \in \text{set}} \overset{\text{pmf part}}{p(x,y)} \right) dx$$

Conditioning:

• for events $A \subseteq B$, suppose $P(B) \neq 0$

then $P(A|B) \triangleq \frac{P(A \cap B)}{P(B)}$
 normalization

event $E \subseteq \Omega$
 $\mathbb{E} \subseteq \mathcal{E}$



$P(B) = \sum_{A \in \text{partition of } \Omega} P(A \cap B)$

for discrete R.V.:

define "conditional pmf"

$p(x|y) \triangleq P(X=x | Y=y) \triangleq \frac{P(\{X=x, Y=y\})}{P(\{Y=y\})}$
 } joint pmf
 } marginal pmf

$p(x|y) \propto p(x,y)$
 "proportional to"
 Normalization cst. $\sum_x p(x,y) = p(y)$

for cts. R.V.: "conditional pdf" $p(x|y) \triangleq \frac{p(x,y)}{p(y)}$

subtle point: $p(x|y)$ is undefined when $p(y)=0$

see Borel-Kolmogorov paradox for weird properties of cond. proba

[Borel-Kolmogorov paradox on wikipedia](#)

* note: independent $X \perp\!\!\!\perp Y$
$$p(x|y) = \frac{p(x,y)}{p(y)} \stackrel{\text{by } \perp}{=} \frac{p(x) \cdot p(y)}{p(y)} = p(x)$$

Bayes rule:

→ invert conditioning

$$p(x|y) = \frac{p(y|x) p(x)}{p(y)}$$

by def.: $p(y|x) = \frac{p(x,y)}{p(x)}$

$$\Rightarrow p(x,y) = p(y|x) p(x) = p(x|y) p(y)$$

↑
"product rule"

example of conditioning: $P(\text{having cancer} | \text{tumor measurement} = "n")$

chain rule: → successive application of product rule

$$p(x_1, \dots, x_n) = p(x_n | x_{1:n-1}) p(x_{1:n-1})$$

$$= \dots = \underbrace{p(x_{n-1} | x_{1:n-2})}_{\text{product rule}} p(x_{1:n-2})$$

$$p(x_{1:n}) = \prod_{i=1}^n p(x_i | x_{1:i-1})$$

always true

convention $1:0 = \emptyset$ $p(x_1 | x_\emptyset) = p(x_1)$

[can also choose any of $n!$ permutations]

→ can be simplified by making conditional independence assumptions
[from directed graphical model]
 $p(x_i | x_{1:i-1}) = p(x_i | x_{\text{pa}(i)})$

conditional independence:

X is cond. indep. of Y given Z notation: $X \perp\!\!\!\perp Y | Z$

$$\Leftrightarrow p(x,y|z) = p(x|z) p(y|z) \quad \forall x \in \Omega_x$$

$$\forall y \in \Omega_y$$

$$\forall z \in \Omega_z \text{ s.t. } p(z) > 0$$

exercise to the reader: prove that $X \perp\!\!\!\perp Y | Z \Rightarrow p(x|y,z) = p(x|z)$

[conditional analog of $X \perp\!\!\!\perp Y \Rightarrow p(z|y) = p(z)$]

example: Z = indicator whether mother carries a genetic disease

X = son 1 has disease
 Y = son 2 " " "

$X \perp\!\!\!\perp Y | Z$



here, X is not "marginally independent" of Y

(i.e. $X \not\perp\!\!\!\perp Y$) $\exists x,u \text{ s.t. } p(x,u) \neq p(x)p(u)$



here, x is not 'marginally independent' of y
 (i.e. $X \not\perp Y$) $\exists x, y$ st. $p(x, y) \neq p(x)p(y)$

$$p(x, y, z) = p(x|z)p(y|z)p(z)$$

16/05

example of pairwise indep. R.V. that are not mutually indep

$X =$ coin flip as $\{0, 1\}$

$Y =$ another indep. coin flip

$$X \perp Y$$

define $Z \triangleq X \text{ xor } Y$

$$Z \perp X$$

$$Z \perp Y$$

$$Z \not\perp (X, Y)$$

other notation :

expectation/mean of a R.V.

$$\mathbb{E}[X] \triangleq \sum_{x \in \Omega_x} x p(x) \quad (\text{for discrete R.V.})$$

$$\int_{\Omega} x p(x) dx \quad (\text{"cts."})$$

$\mathbb{E}[\cdot]$ is a linear operator $\Rightarrow \mathbb{E}[aX + Y] = a\mathbb{E}[X] + \mathbb{E}[Y]$
 a is a scalar

Variance :

$$\text{Var}[X] \triangleq \mathbb{E}[(X - \mathbb{E}[X])^2]$$

$$= \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

"measure of dispersion"



Cumulative distribution fun. CDF of X : $F_X(t) = \mathbb{P}\{X \leq t\} = \int_{-\infty}^t p(x) dx$

$$\left. \frac{d}{dt} F_X(t) \right|_{t=x} = p(x) \quad \text{for a cts. R.V.}$$

(for a cts. R.V.)

parametric models :

\hookrightarrow a family of distributions

$$\mathcal{P}_{\Theta} = \{p(\cdot | \theta) | \theta \in \Theta\}$$

possible pdf's / pdf's depend on parameter θ

from context

"support of distribution" i.e. Ω_X is usually fixed for all θ (but not always)

examples: $\Omega_X = \{0, 1\}$ for coin flip

abuse of notation: $\{p(x; \theta) | \theta \in \Theta\}$

examples: $X \sim \text{Bern}(\theta)$ for coin flip

abuse of notation: $\{p(x; \theta) \mid \theta \in \Theta\}$

'better notation' $\{p(X; \theta) \mid \theta \in \Theta\}$

$\Omega_X = [0, 1]$
for gamma dist. family

notation: $X \sim \text{Bern}(\theta)$

"R.V. X is distributed as a Bernoulli dist. with parameter θ "

$$p(x; \theta) = \text{Bern}(x; \theta) = \theta^x (1-\theta)^{1-x} \quad x \in \{0, 1\}$$

variable for pmf

Bernoulli: prob. of a coin flip $\Omega_X = \{0, 1\}$

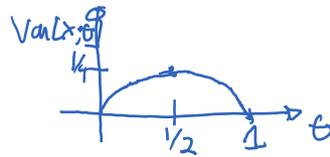
$$\Theta = [0, 1]$$

$$p(x=1) = \theta$$

Bayesian notation $P\{X=1 \mid \theta\} = \theta$

$$E[X] = \theta$$

$$\text{Var}[X] = \theta(1-\theta)$$



another example:

$$p(x; (\mu, \sigma^2)) = N(x; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

or

$$N(x; \mu, \sigma^2)$$

binomial dist.:

model n indep. coin flips

sum of n indep. $\text{Bern}(\theta)$ R.V.

let $X_i \stackrel{\text{iid}}{\sim} \text{Bern}(\theta) \rightarrow$ (mutually) independent and identically distributed

implicitly talking $\leftrightarrow X_1, \dots, X_n$

let $X = \sum_{i=1}^n X_i$ then we have $X \sim \text{Bin}(n, \theta)$

"binomial dist. with parameter n & θ "

$$\Omega_X = \{0, 1, \dots, n\}$$

$$\text{Bern}(x_i; \theta) = \theta^{x_i} (1-\theta)^{1-x_i}$$

$$\text{pmf: } p(x; \theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x}$$

$$\prod_{i=1}^n \text{Bern}(x_i; \theta) = \theta^{\sum_{i=1}^n x_i} (1-\theta)^{n - \sum_{i=1}^n x_i}$$

of ways to choose x elements out of n
"n choose x"

$$\binom{n}{x} \triangleq \frac{n!}{x!(n-x)!}$$

mean: $X = \sum_1^n x_i$

$E[X] = \sum_1^n E[x_i] = n\theta$

Similarly, $Var(\sum_1^n x_i) \stackrel{\text{indep.}}{=} \sum_1^n Var(x_i) = n\theta(1-\theta)$

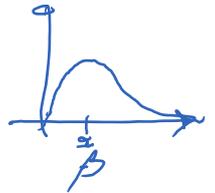
Other distributions: Poisson (λ) $\Omega_X = \{0, 1, \dots\} = \mathbb{N}$ [count data]

Gaussian in 1d $N(\mu, \sigma^2)$ $\Omega_X = \mathbb{R}$
mean, variance

Gamma $\Gamma(\alpha, \beta)$ $\Omega_X = \mathbb{R}_+$

shape inverse aka rate

mean $\propto \beta$ variance $\propto \beta^2$



Other: Laplace, Cauchy, exponential dist. $\hookrightarrow \Gamma(1, \beta)$, beta is a Dirichlet on 2 elements $\hookrightarrow \Omega_X = [0, 1]$

Dirichlet